

SOME RESULTS FROM MULTIVARIATE STATISTICS

Covariance matrices

Let \mathbf{x} be an $n \times 1$ random vector. The *covariance matrix* of \mathbf{x} is an $n \times n$ matrix, denoted $\text{Cov}(\mathbf{x})$, whose (i, j) th entry is the covariance between x_i and x_j . An equivalent definition is

$$\text{Cov}(\mathbf{x}) = \text{E}\left[\{\mathbf{x} - \text{E}(\mathbf{x})\}\{\mathbf{x} - \text{E}(\mathbf{x})\}^T\right].$$

If \mathbf{x} is a random vector, \mathbf{A} is a constant matrix and \mathbf{c} is a constant vector whose dimensions are such that $\mathbf{Ax} + \mathbf{c}$ is defined then

$$\text{E}(\mathbf{Ax} + \mathbf{c}) = \mathbf{A}\text{E}(\mathbf{x}) + \mathbf{c}$$

and

$$\text{Cov}(\mathbf{Ax} + \mathbf{c}) = \mathbf{A}\text{Cov}(\mathbf{x})\mathbf{A}^T.$$

The following relationships hold between conditional and unconditional means and covariance matrices:

$$\begin{aligned}\text{E}(\mathbf{y}) &= \text{E}\{\text{E}(\mathbf{y}|\mathbf{x})\} \\ \text{Cov}(\mathbf{y}) &= \text{E}\{\text{Cov}(\mathbf{y}|\mathbf{x})\} + \text{Cov}\{\text{E}(\mathbf{y}|\mathbf{x})\}.\end{aligned}$$

Multivariate normal distribution

The $d \times 1$ random vector \mathbf{x} has a multivariate normal distribution with mean $\boldsymbol{\mu}$ and invertible covariance matrix $\boldsymbol{\Sigma}$ if its probability density function is

$$[\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}] = (2\pi)^{-d/2} |\boldsymbol{\Sigma}|^{-1/2} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right\}$$

and denote this by

$$\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}).$$

If

$$\begin{bmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \end{bmatrix} \sim N\left(\begin{bmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{bmatrix}\right)$$

is a general partitioned normal random vector, then the marginal distribution of \mathbf{x}_2 is $N(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})$ and the conditional distribution of \mathbf{x}_2 given \mathbf{x}_1 is

$$[\mathbf{x}_2|\mathbf{x}_1] \sim N(\boldsymbol{\mu}_2 + \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}(\mathbf{x}_1 - \boldsymbol{\mu}_1), \boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12}).$$

Analogous results hold for \mathbf{x}_1 . Alternatively, $\mathbf{x}_2|\mathbf{x}_1$ may be expressed in terms of the submatrices of the *inverse* covariance matrix of $[\mathbf{x}_1^T, \mathbf{x}_2^T]^T$. If

$$\begin{bmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \end{bmatrix} \sim N\left(\begin{bmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{bmatrix}, \begin{bmatrix} \mathbf{Q}_{11} & \mathbf{Q}_{12} \\ \mathbf{Q}_{21} & \mathbf{Q}_{22} \end{bmatrix}^{-1}\right)$$

then

$$[\mathbf{x}_2|\mathbf{x}_1] \sim N(\boldsymbol{\mu}_2 - \mathbf{Q}_{22}^{-1}\mathbf{Q}_{21}(\mathbf{x}_1 - \boldsymbol{\mu}_1), \mathbf{Q}_{22}^{-1}).$$